

MATH4210: Financial Mathematics Tutorial 5

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Convergence of r.v.s

Let $(\Omega, \mathcal{F}, \mathbb{P})$ be a probability space. X and $\{X_n\}$ are \mathbb{R} valued (sequence of) r.v.s.

Definition (Convergence almost surely)

Denote by $X_n \rightarrow X$ a.s. (almost surely) if

$$\mathbb{P}[\{\omega \in \Omega : \lim_{n \rightarrow \infty} X_n(\omega) = X(\omega)\}] = 1$$

Definition (Convergence in Probability)

Denote by $X_n \rightarrow X$ in probability if for any $\rho > 0$

$$\lim_{n \rightarrow \infty} \mathbb{P}[\{\omega \in \Omega : |X_n(\omega) - X(\omega)| \geq \rho\}] = 0$$

Convergence of r.v.s

Proposition *proved*

$X_n \rightarrow X$ a.s. implies $X_n \rightarrow X$ in probability.

Proposition *admit*

*(Search
Boel-Cantelli Lemma)*

$X_n \rightarrow X$ in probability implies there exists a subsequence of X_n converging to X a.s..

Definition (Convergence in Law (in Distribution))

Let F_n and F be the c.d.f. of X_n and X for all $n \in \mathbb{N}$. $X_n \rightarrow X$ in Law (in Distribution) if

$$\lim_{n \rightarrow \infty} F_n(x) = F(x)$$

for any $x \in \mathbb{R}$ where F is continuous at x .

Proposition *admit*

$X_n \rightarrow X$ in probability implies $X_n \rightarrow X$ in Law.

Convergence of r.v.s.

Definition

Given $p > 0$, denote by $X_n \rightarrow X$ in L^p if

$$\lim_{n \rightarrow \infty} \mathbb{E}[|X_n - X|^p] = 0$$

$$\lim_{n \rightarrow \infty} \int_{\mathbb{R}} |f_n - f|^p(x) dx = 0$$

Question

- (a). Show that $X_n \rightarrow X$ in L^2 implies $X_n \rightarrow X$ in L^1 .
(b). Show that $X_n \rightarrow X$ in L^p implies $X_n \rightarrow X$ in probability.

(a). proof : $(\lim_{n \rightarrow \infty} \mathbb{E}[|X_n - X|^2] = 0) \Rightarrow (\lim_{n \rightarrow \infty} \mathbb{E}[|X_n - X|] = 0)$.

Assume \leftarrow is true.

Fix $n \in \mathbb{N}$

$$E[|X_n - x|] \leq \sqrt{E[|X_n - x|^2]}$$

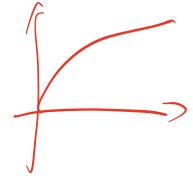
$$E[|X_n - x|] = E[\sqrt{|X_n - x|^2}]$$

$$\leq \sqrt{E[|X_n - x|^2]} \quad \leftarrow \text{Jensen's inequality.}$$

$\rightarrow 0.$

Since $x \mapsto \sqrt{x}$ is continuous and $\lim_{n \rightarrow \infty} E[|X_n - x|^2] = 0$ □

if f is convex,
then $f(E(x)) \leq E(f(x))$
if f is concave
then:
 $E(f(x)) \leq f(E(x))$



(b). Proof:

Fix $\rho > 0$. fix $n \in \mathbb{N}$.

$$P[\{\omega \in \Omega : |X_n - x| \geq \rho\}] \xrightarrow{n \rightarrow \infty} 0$$

$$= P[|X_n - x|^p \geq \rho^p]$$

$$\leq \frac{1}{\rho^p} E[|X_n - x|^p] \rightarrow \text{by Markov's inequality}$$

$$P[|X| > \eta] \leq \frac{1}{\eta} E[|X|]$$

$\rightarrow 0$ as $n \rightarrow \infty$.

by L^p convergence.

□

$$X_n \xrightarrow{\text{a.s.}} X$$

Conditions

Montone CV theorem
Dominated CV theorem

↓ subsequence

$$X_n \rightarrow X \text{ in } L^p.$$

$$X_n \xrightarrow{P} X$$

↓

$$X_n \xrightarrow{L^1} X$$

Brownian Motions

Question

Let $f : \mathbb{R} \rightarrow \mathbb{R}$ be such that for some constant $C > 0$, $f(x) < e^{C|x|}$ for all $x \in \mathbb{R}$. Define

$$u(t, x) = \mathbb{E}[f(B_T) | B_t = x] = \mathbb{E}[f(B_T - B_t + x)].$$

Show that

(a)

$$\partial_x u(t, x) = \mathbb{E}\left[\frac{B_T - B_t}{T - t} f(B_T - B_t + x)\right]$$

(b)

$$\partial_x^2 u(t, x) = \mathbb{E}\left[\frac{(B_T - B_t)^2 + (T - t)}{(T - t)^2} f(B_T - B_t + x)\right]$$

$$= \mathbb{E}\left[\frac{Z^2 + (T-t)}{(T-t)^2} f(Z+x)\right], \quad Z \sim N(0, T-t)$$

(a). Proof: Fix $(t, x) \in \mathbb{R}_+ \times \mathbb{R}$.

$$u(t, x) = \mathbb{E}[f(B_T - B_t + x)]$$

Denote $Z = N(0, T-t)$, denote by p_Z its pdf.

$$\text{So } u(t, x) = \mathbb{E}[f(Z+x)].$$

$$= \int_{\mathbb{R}} f(y+x) p_Z(y) dy.$$

$$\text{Then } \partial_x u(t, x) = \int_{\mathbb{R}} f'(y+x) p_Z(y) dy.$$

$$\text{IBP. } = \underbrace{\left[f(y+x) p_Z(y) \right]_{y \rightarrow -\infty}^{y \rightarrow +\infty}}_{\downarrow 0} - \int_{\mathbb{R}} f(y+x) p_Z'(y) dy.$$

$$p_Z'(y) = \frac{1}{\sqrt{2\pi(T-t)}} e^{-\frac{y^2}{2(T-t)}} \cdot \left(-\frac{y}{(T-t)} \right)$$

$$\text{So } \partial_x u(t, x) = 0 + \frac{1}{T-t} \int_{\mathbb{R}} y f(y+x) p_Z(y) dy.$$

$$= \mathbb{E}\left[\frac{f(Z+x)}{T-t} \cdot Z \right]$$

$$= \mathbb{E}\left[\frac{f(B_T - B_t + x)}{T-t} (B_T - B_t) \right]$$

□

(b). Similarly. We apply the IBP

Greeks of Option

$$d_1: x \mapsto \frac{\ln(x/K) + (r + \sigma^2/2)(T-t)}{\sigma\sqrt{T-t}}$$

$$C_E: (t, x) \mapsto xN(d_1(x)) - e^{-r(T-t)}KN(d_2(x))$$

Question

Consider the European call option price at time t :

$$C_E(t, S_t) = S_t N(d_1) - e^{-r(T-t)}KN(d_2)$$

where $d_1 = \frac{\ln(S_t/K) + (r + \sigma^2/2)(T-t)}{\sigma\sqrt{T-t}}$ and $d_2 = d_1 - \sigma\sqrt{T-t}$. Compute

(a) Delta: $\Delta = \partial_x C_E(t, S_t)$.

(b) Gamma: $\Gamma = \partial_x^2 C_E(t, S_t)$.

$$(a). d_1(S_t) = \frac{\ln(S_t/K) + (r + \sigma^2/2) \cdot \tau}{\sigma\sqrt{\tau}}, \quad \tau = T-t$$

$$\partial_x d_1(S_t) = \frac{1}{\sigma\sqrt{\tau}} \cdot \frac{1}{S_t} = \partial_x d_2(S_t)$$

$$N(d_1(S_t)) = \int_{-\infty}^{d_1(S_t)} f_Z(y) dy, \text{ where } Z \sim N(0,1) \\ f_Z \text{ its pdf.}$$

$$\text{So. } \frac{\partial}{\partial x} N(d_1(S_t)) = f_Z(d_1(S_t)) \cdot \partial_x d_1(S_t).$$

Therefore by chain rules

$$\begin{aligned} \partial_x C_E(t, S_t) &= N(d_1(S_t)) + S_t \cdot f_Z(d_1(S_t)) \cdot \partial_x d_1(S_t) \\ &\quad - e^{r\tau} k \cdot f_Z(d_2(S_t)) \cdot \partial_x d_2(S_t) \end{aligned}$$

$$= N(d_1(S_t)) + \partial_x d_1(S_t) \cdot (S_t f_Z(d_1(S_t)) - e^{r\tau} k f_Z(d_2(S_t)))$$

$$= N(d_1(S_t)) + \partial_x d_1(S_t) \cdot \frac{1}{\sqrt{2\pi}} \left(S_t \cdot e^{-\frac{d_1^2}{2\sigma^2}} - e^{r\tau} k e^{-\frac{d_2^2}{2\sigma^2}} \right)$$

Note that.

$$d_1^2 - d_2^2 = (d_1 + d_2)(d_1 - d_2)$$

$$= \left(2 \frac{\ln(S_t/k) + (r + \frac{\sigma^2}{2})\tau}{\sigma \sqrt{\tau}} - \sigma \sqrt{\tau} \right) \cdot (\sigma \sqrt{\tau})$$

$$= \underline{2 \ln(S_t/k) + 2r\tau}$$

$$\text{So } \partial_x C_E(t, S_t) = N(d_1(S_t)) + \frac{\partial_x d_1(S_t)}{\sqrt{2\pi}} e^{-\frac{d_1^2}{2\sigma^2}} \left(S_t - e^{r\tau} k \cdot e^{-\frac{d_2^2 - d_1^2}{2\sigma^2}} \right)$$

$$S_t - e^{-r\tau} k \cdot e^{-\frac{d_2^2 - d_1^2}{2\sigma^2}} = 0$$

(A) a consequence

$$\partial_x C_E(t, S_t) = N(d_1). \quad \square$$

$$(b). \Gamma = \partial_x^2 C_E(t, S_t)$$

$$= \partial_{xx} N(d_1(S_t))$$

$$= f_z(d_1(S_t)) \cdot \partial_x d_1(S_t).$$

$$= \frac{1}{\sigma \sqrt{t} S_t} f_z(d_1(S_t)).$$